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## META-ANALYSIS

# Probiotics for the prevention of gestational diabetes mellitus: A meta-analysis of randomized controlled trials

Xue Li<sup>1</sup>, Luwen Zhang<sup>2</sup>, Yuanqi He<sup>1</sup>, Dandan Zhang<sup>1</sup>, and Shihong Zhang<sup>3</sup>

Changes in intestinal microbiota have been shown to be involved in the development of gestational diabetes mellitus (GDM).

We performed a meta-analysis to systematically evaluate the potential role of probiotics in the prevention of GDM. A systematic literature search was performed in electronic databases, including PubMed, Cochrane Library, Embase, Web of Science, Wanfang, and China National Knowledge Infrastructure (CNKI) to obtain relevant randomized controlled studies. A random-effects model was used to pool the results by incorporating the impact of the potential heterogeneity. Meta-regression and subgroup analyses were conducted to evaluate the source of heterogeneity. Fourteen studies involving 3527 pregnant women were included. Results showed that probiotics significantly reduced the incidence of GDM as compared to control (risk ratio [RR]: 0.71, 95% confidence interval [CI]: 0.52–0.96,  $P = 0.03$ ) with significant heterogeneity ( $I^2 = 73\%$ ). The meta-regression showed that the body mass index (BMI) of women was positively associated with the RR for the effect of probiotics on GDM (coefficient = 0.084,  $P = 0.01$ ). The results of subgroup analyses also suggested that probiotics significantly reduced the risk of GDM in women with BMI < 26 kg/m<sup>2</sup>, but not in those with BMI ≥ 26 kg/m<sup>2</sup> ( $P$  for subgroup difference = 0.001). In addition, the preventative efficacy of probiotics on GDM was remarkable in women < 30 years, but not in those ≥ 30 years ( $P$  for subgroup difference < 0.001). In conclusion, probiotics may be effective in reducing the risk of GDM, particularly for women with lower BMI and younger age.

**Keywords:** Gestational diabetes mellitus (GDM), probiotics, prevention, incidence, meta-analysis.

## Introduction

Gestational diabetes mellitus (GDM) is a prevalent metabolic disorder that occurs during pregnancy [1, 2]. Existing literature suggests that the prevalence of GDM among pregnant individuals ranges from 15% to 20% [1]. Risk factors associated with GDM include advanced maternal age, elevated body mass index (BMI), familial history of type 2 diabetes mellitus (T2DM), and a prior history of GDM in a previous pregnancy [3]. Emerging research indicates that GDM is not only linked to immediate adverse outcomes, such as miscarriage, preterm birth, and macrosomia [4, 5], but it is also associated with a range of long-term health risks for both mothers and their offspring, including maternal and child obesity, increased risk of type 2 diabetes, and heightened maternal susceptibility to cancer and cardiovascular diseases [4, 6, 7]. Consequently, there is a pressing need for the development of innovative approaches to prevent the onset of GDM [8].

Pregnancy has been associated with disruptions in the homeostasis of intestinal microbiota, with a notable increase in actinobacteria and proteobacteria observed in 60%–70% of women [9, 10]. Studies have shown that women with GDM exhibit more pronounced alterations in gut microbiota compared to those without GDM, resembling patterns seen in

non-pregnant women with T2DM [11, 12]. This suggests a potential role of gut microbiota in the development of GDM. Probiotics, as living microorganisms, play a beneficial role in restoring and maintaining the balance of gut microbiota composition [13]. In T2DM patients, the use of probiotics has been linked to a reduction in insulin resistance and enhancement of glycemic control [14, 15]. Furthermore, in females with a confirmed diagnosis of GDM, supplementation with probiotics has demonstrated improvements in hyperglycemia and dyslipidemia, as well as a decrease in the birth weight of their offspring [16–18]. Similarly, probiotics supplementation has been suggested to improve glycemic control via multiple mechanisms, such as reducing inflammation, enhancing the production of short-chain fatty acids (SCFAs), regulation of gut microbiota, improving insulin sensitivity, and preventing excessive weight gain [19, 20]. However, conflicting findings arise from previous studies examining the efficacy of probiotics in preventing GDM [21]. Two meta-analyses conducted previously did not find significant evidence to support the use of probiotics in reducing the risk of GDM [22, 23]. However, they included only five to six studies and significant heterogeneity, which were not explored due to the limited number of available studies, was observed in both [22, 23]. Additional

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57 randomized controlled trials have been published since [24–29].  
58 Accordingly, the aim of our study was to perform an updated  
59 meta-analysis to comprehensively evaluate the influence of  
60 probiotics supplementation on the incidence of GDM in preg-  
61 nant women.

## 62 Materials and methods

63 This study is in accordance with the guidelines of Preferred  
64 reporting items for systematic reviews and meta-analyses  
65 (PRISMA) [30, 31] and the Cochrane Handbook [32].

### 66 Study inclusion and exclusion criteria

67 The principle of PICOS, which is explained below, was utilized  
68 to determine the inclusion criteria for the meta-analysis.

69 P (participants): Women planning to conceive or at early  
70 pregnancy; I (intervention): Probiotics supplements during  
71 pregnancy, with no restrictions to the strains, timing, or dose  
72 of probiotics; C (control): Placebo or no additional treatment;  
73 O (outcomes): Reported the incidence of GDM during follow-  
74 up. The methods and criteria for the diagnosis of GDM were in  
75 accordance with those reported in the original studies. S (study  
76 design): Only RCTs with parallel groups that were published as  
77 complete articles in English or Chinese in peer-reviewed jour-  
78 nals were deemed eligible for study design. Non-randomized  
79 studies, studies not including women planning to conceive or at  
80 early pregnancy, not with an intervention of probiotic supple-  
81 mentation, or not reporting the outcome of GDM incidence were  
82 excluded. In case studies with potentially overlapping patient  
83 populations were found, the meta-analysis included the one  
84 that had the larger sample size.

### 85 Literature search strategy

86 To identify studies in Medline (PubMed), CENTER (Cochrane  
87 Library), Embase (Ovid), Web of Science, Wanfang, and China  
88 National Knowledge Infrastructure (CNKI), a search strategy  
89 was employed that encompassed the following criteria by a  
90 combination of the keywords: (1) “probiotic” OR “probiotics”  
91 OR “lactobacillus” OR “lactobacilli” OR “bifidobacteria” OR “bifi-  
92 dobacterium”; (2) “gestational diabetes mellitus” OR “GDM”  
93 OR [(“gestational” OR “pregnancy” OR “pregnant”) AND (“dia-  
94 betes” OR “diabetic” OR “hyperglycemia”)]; and (3) “random”  
95 OR “randomized” OR “randomized” OR “randomly” OR “allo-  
96 cated” OR “control” OR “placebo.” Our focus was solely on  
97 research that involved human participants. In addition, we con-  
98 ducted a manual search for references to relevant reviews and  
99 primary articles. The most recent database search was con-  
100 ducted on December 21, 2023.

### 101 Extraction of data and assessment of study quality

102 Two authors conducted separate searches in databases, gath-  
103 ered information, and assessed the quality. In case of any dis-  
104 agreements, the corresponding author was consulted to reach  
105 a consensus. For the study, various data were gathered includ-  
106 ing general details, characteristics of the study design, par-  
107 ticipant characteristics, age, BMI, proportions of women with  
108 primipara, use of lifestyle recommendations (diet and exercise),  
109 details of interventions (probiotics used, timing, and dose),

regimens of controls, and criteria for the diagnosis of GDM. 110  
Cochrane’s Risk of Bias Tool [32] was used to evaluate the qual- 111  
ity of RCTs included in this review. It assessed seven domains, 112  
including the generation of random sequence, concealment of 113  
allocations, blinding of participants and personnel, blinding 114  
of outcome evaluation, incomplete result data, and selective 115  
reporting of outcomes. 116

### 117 Statistical analysis

118 The incidence of GDM, compared between women with pro- 119  
biotics supplementation and women in the control group, was 120  
summarized as risk ratio (RR) and corresponding 95% confi- 121  
dence interval (CI). The outcome data was extracted using the 122  
intention-to-treat principle. The Cochrane Q test was used to 123  
investigate the heterogeneity among the included studies [32]. 124  
Furthermore, the  $I^2$  statistic was calculated, where  $I^2 > 50\%$  125  
suggested statistical heterogeneity [33]. To incorporate poten- 126  
tial heterogeneity, a random-effect model was employed for 127  
pooling the data [32]. For outcomes of adequate datasets (10 or 128  
above), meta-regression and subgroup analyses according to 129  
study characteristics were performed to evaluate the source of 130  
heterogeneity. The meta-regression analysis tested the signifi- 131  
cance of the individual study characteristics’ influence on the 132  
results of the meta-analysis, with a  $P$  value  $< 0.05$  indicating 133  
a significant modification effect. A positive coefficient demon- 134  
strated that the evaluated study characteristics are positively 135  
related to the OR of the results, while a negative coefficient 136  
demonstrated that the evaluated study characteristics are neg- 137  
atively related to the OR of the results. These characteristics 138  
included study country, mean age, BMI, timing, and dose of 139  
probiotics supplementation, and the risk of GDM of the studied 140  
females as reflected by the incidence of GDM in the control 141  
groups. Medians of continuous variables were selected as the 142  
cutoffs to define the subgroups. Publication bias was evalu- 143  
ated using Egger’s test for regression asymmetry and funnel 144  
plots [34]. A  $P$  value  $< 0.05$  suggested a statistically signifi- 145  
cant distinction. The statistical analysis was conducted using 146  
RevMan (Version 5.1; Cochrane, Oxford, UK) and Stata (Version 147  
12.0; Stata Corporation, USA) software.

## 148 Results

### 149 Literature search

150 The process of acquiring literature is illustrated in Figure 1. In 151  
summary, a total of 719 articles were obtained through database 152  
searches, with 530 remaining after removing duplicates. A total 153  
of 494 articles were subsequently excluded by screening via 154  
titles and abstracts, primarily because they were not relevant 155  
to the objective of the study. After reading the full text, an 156  
additional 22 articles out of the initial 36 were excluded due to 157  
the reasons outlined in Figure 1. At last, 14 RCTs [24–29, 35–42] 158  
were available for the subsequent meta-analysis. 159

### 160 Study characteristics and data quality evaluation

161 Table 1 provides a summary of the studies included in the 162  
meta-analysis. In total, there were 14 RCTs involving 3527 163  
females who were planning to conceive in the upcoming six 164  
months or at early pregnancy [24–29, 35–42]. These studies

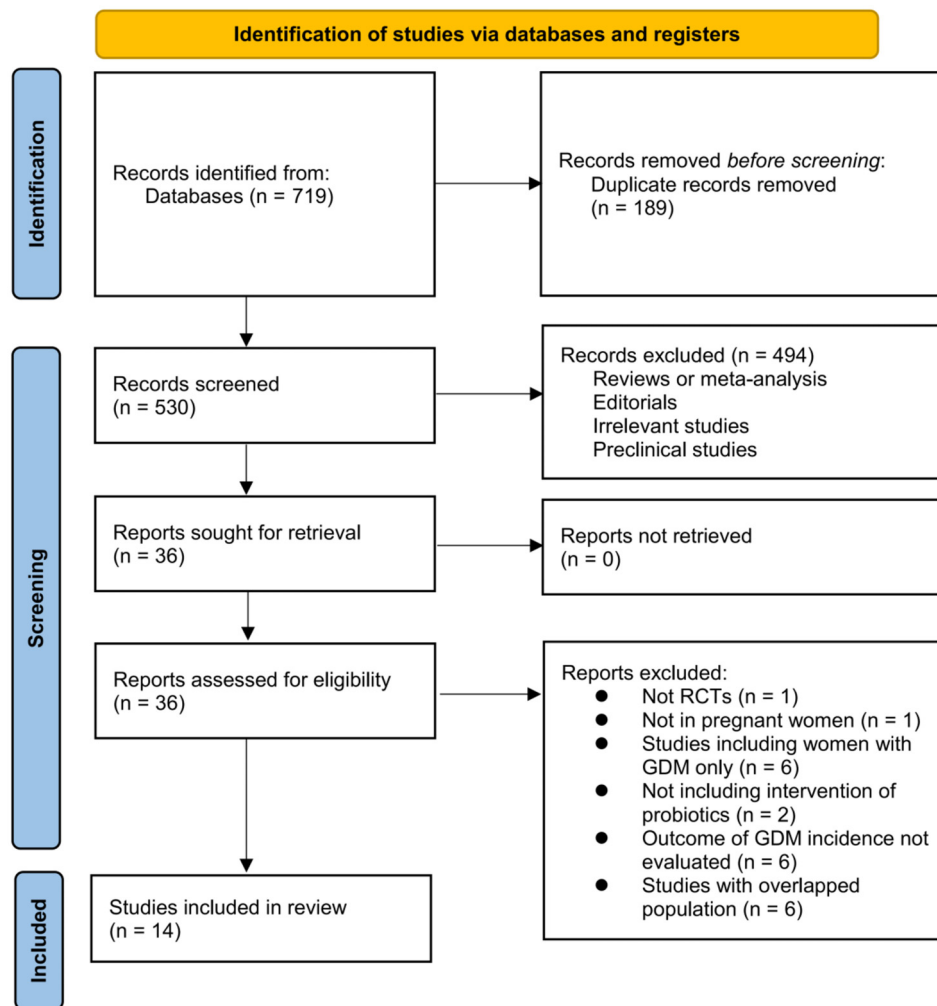


Figure 1. Flowchart of the literature search. GDM: Gestational diabetes mellitus.

164 were published between 2010 and 2022 and carried out in  
 165 Finland, Ireland, New Zealand, Australia, China, Denmark,  
 166 Iran, the United Kingdom, Singapore, and Pakistan. The mean  
 167 ages of the females were 27–34 years, and the mean BMI scores  
 168 were 21–39 kg/m<sup>2</sup>. The proportions of females with primipara  
 169 varied between 15.0%–63.5%. In four studies, dietary recom-  
 170 mendation was also provided to females of the intervention  
 171 and control groups [25, 35, 36, 41]. However, no evaluation has  
 172 been performed regarding the diet or physical activities pre-  
 173 and post-intervention among these studies. Multiple differ-  
 174 ent strains were used for probiotics supplementation, such as  
 175 *Lactobacillus rhamnosus*, *Lactobacillus acidophilus*, *Lactobacillus*  
 176 *salivarius*, *Bifidobacterium lactis*, *Bifidobacterium longum*, *Bifi-*  
 177 *dobacterium bifidum*, *Lactobacillus casei*, *Lactobacillus bulgaricus*,  
 178 *Lactobacillus plantarum*, *Lactobacillus paracasei*, *Bifidobacterium*  
 179 *breve*, *Bifidobacterium infantis*, and *Streptococcus thermophilus*,  
 180 with *Lactobacillus rhamnosus* GG as the most commonly used  
 181 probiotics strain. Most of the included studies used multi-  
 182 ple strains as intervention except for three studies [36–38],  
 183 in which single-strain probiotics were used. The timing for  
 184 the starting of probiotics supplementation varied among the  
 185 included studies, ranging from within the first trimester to the

186 gestational age (GA) of 24 weeks. The total doses of probiotics  
 187 were 1–50 × 10<sup>9</sup> colony-forming units per day. As for the con-  
 188 trols, placebo capsules were used in 12 studies [26–29, 35–42],  
 189 while for the other two studies, no additional treatment was  
 190 considered as controls [24, 25]. The incidence of GDM was  
 191 diagnosed with the International Association of Diabetes in  
 192 Pregnancy Study Group criteria [43] in all the studies using a  
 193 “one-step” 2-h 75 g oral glucose tolerance test (OGTT) except  
 194 for one study [36], in which GDM was diagnosed with the  
 195 American College of Obstetricians and Gynecologists criteria  
 196 using a “two-step” 3-h 100 g OGTT test [44]. Compliance data  
 197 were reported in three studies, with similar mean adherence  
 198 rates of 94.5% [37], 88.4% [40], and >90% [38] between females  
 199 of the intervention and control groups, indicating good com-  
 200 pliance. The incidence of adverse events was reported in two  
 201 studies [26, 40]. Only mild discomfort related to the treatments  
 202 was reported, which was similar in females in the intervention  
 203 and control groups, with gastrointestinal symptoms being the  
 204 most common symptoms.

205 Table 2 provides a detailed analysis of the included RCTs  
 206 using Cochrane’s Risk of Bias Tool. One of the included studies  
 207 was open-label [25], another one was single-blinded [24], while

Table 1. Characteristics of the included studies

| Study           | Location    | Design    | Participants   | Patient number | Mean age (years) | Mean BMI (kg/m <sup>2</sup> ) | Primipara (%) | Lifestyle recommendations | Timing of intervention                         | Intervention   | Total dose (10 <sup>9</sup> cfu/d) | Control                 | GDM diagnosis   |
|-----------------|-------------|-----------|--|----------------|------------------|-------------------------------|---------------|---------------------------|--|--|------------------------------------|-------------------------|-----------------|
| Luoto 2010      | Finland     | R, DB, PC | Women at early pregnancy with no chronic metabolic diseases                            | 152            | 29.9             | 23.6                          | 57.9          | Diet only                 | First trimester to delivery                    | <i>Lactobacillus rhamnosus</i> GG and <i>Bifidobacterium lactis</i> Bb12   | 20                                 | Placebo capsule         | IADPSG criteria |
| Lindsay 2014    | Ireland     | R, DB, PC | Obese women at early pregnancy   | 138            | 31.2             | 33.6                          | 44.9          | Diet only                 | Second trimester (GA: 24 weeks) to delivery    | <i>Lactobacillus salivarius</i> UCC118   | 1                                  | Placebo capsule         | ACOG criteria   |
| Wickens 2017    | New Zealand | R, DB, PC | Pregnant women with a personal or partner history of atopic disease at early pregnancy | 373            | 34               | 25.5                          | NR            | NR                        | Second trimester (GA: 14~16 weeks) to delivery | <i>Lactobacillus rhamnosus</i> HN001   | 6                                  | Placebo capsule         | IADPSG criteria |
| Okesene 2019    | New Zealand | R, DB, PC | Women at early pregnancy with BMI > 30 kg/m <sup>2</sup>                               | 230            | 28.7             | 38.6                          | 31.7          | No dietary recommendation | Second trimester (GA: 13~17 weeks) to delivery | <i>Lactobacillus rhamnosus</i> GG and <i>Bifidobacterium lactis</i> Bb12   | 6.5                                | Placebo capsule         | IADPSG criteria |
| Pellonpera 2019 | Finland     | R, DB, PC | Women at early pregnancy with BMI > 25 kg/m <sup>2</sup>                               | 190            | 30.6             | 29.8                          | 47.9          | No dietary recommendation | Second trimester (GA: 18 weeks) to delivery    | <i>Lactobacillus rhamnosus</i> HN001 and <i>Bifidobacterium animalis</i> ssp. <i>lactis</i> 420  | 20                                 | Placebo capsule         | IADPSG criteria |
| Callaway 2019   | Australia   | R, DB, PC | Women at early pregnancy with BMI > 25 kg/m <sup>2</sup>                               | 411            | 31.5             | 31.8                          | 38.7          | NR                        | Second trimester (GA: 20 weeks) to delivery    | <i>Lactobacillus rhamnosus</i> GG and <i>Bifidobacterium lactis</i> Bb12   | 2                                  | Placebo capsule         | IADPSG criteria |
| Wang 2019       | China       | R, SB     | Women at early pregnancy   | 400            | 27.2             | 21.5                          | NR            | NR                        | Second trimester (GA: 16 weeks) to delivery    | <i>Lactobacillus rhamnosus</i> GG and <i>Bifidobacterium lactis</i> Bb12   | 2                                  | No additional treatment | IADPSG criteria |
| Halkjar 2020    | Denmark     | R, DB, PC | Obese women at early pregnancy   | 49             | 30.7             | 31.9                          | NR            | NR                        | Second trimester (GA: 14~20 weeks) to delivery | <i>Streptococcus thermophilus</i> DSM 24,731, <i>bifidobacteria</i> and <i>lactobacilli</i>  | 45                                 | Placebo capsule         | IADPSG criteria |
| Cao 2020        | China       | R, OL     | Women at early pregnancy   | 100            | 33.8             | 23.6                          | 50%           | Diet only                 | Second trimester (GA: 13~14 weeks) to delivery | <i>Lactobacillus rhamnosus</i> GG and HN001, <i>Limosilactobacillus reuteri</i> CECT5716, <i>Bifidobacterium animalis</i> ssp. <i>lactis</i> HN019 | 15                                 | No additional treatment | IADPSG criteria |

(Continued)

Table 1. Continued

| Study          | Location                                   | Design    | Participants   | Patient number | Mean age (years) | Mean BMI (kg/m <sup>2</sup> ) | Primipara (%) | Lifestyle recommendations | Timing of intervention                         | Intervention   | Total dose (10 <sup>9</sup> cfu/d) | Control          | GDM diagnosis   |
|----------------|--|-----------|--|----------------|------------------|-------------------------------|---------------|---------------------------|--|--|------------------------------------|------------------|-----------------|
| Asgharian 2020 | Iran                                       | R, DB, PC | Women at early pregnancy with BMI > 25 kg/m <sup>2</sup> | 128            | 29.5             | 29.8                          | 15%           | Diet only                 | Second trimester (GA: 24 weeks) to delivery    | <i>Lactobacillus acidophilus</i> La5 and <i>Bifidobacterium lactis</i> Bb12  | 50                                 | Placebo yoghurts | IADPSG criteria |
| Shahriari 2021 | Iran                                       | R, DB, PC | Women at early pregnancy with high risk for GDM          | 507            | 32               | 30.2                          | NR            | No dietary counseling     | Second trimester (GA: 14 weeks) to delivery    | <i>Lactobacillus acidophilus</i> LA1, <i>Bifidobacterium longum</i> sp54 cs, and <i>Bifidobacterium bifidum</i> sp9 cs | 15                                 | Placebo capsule  | IADPSG criteria |
| Godfrey 2021   | United Kingdom, Singapore, and New Zealand | R, DB, PC | Women planning to conceive in upcoming 6 months          | 577            | 30.3             | 25.7                          | 63.5          | NR                        | Before pregnancy to delivery                   | <i>Lactobacillus rhamnosus</i> NCC 4007 and <i>Bifidobacterium animalis</i> subspecies lactis NCC 2818                 | 2                                  | Placebo capsule  | IADPSG criteria |
| Baloch 2022    | Pakistan                                   | R, DB, PC | Women at early pregnancy with high risk for GDM          | 160            | 30               | 26                            | NR            | NR                        | Second trimester (GA: 13~14 weeks) to delivery | <i>Streptococcus</i> , bifidobacteria and lactobacilli   | 5                                  | Placebo capsule  | IADPSG criteria |
| Liu 2022       | China                                      | R, DB, PC | Women at early pregnancy                                 | 112            | 29.7             | 22.6                          | NR            | NR                        | Second trimester (GA: 20 weeks) to delivery    | <i>Streptococcus</i> , bifidobacteria and lactobacilli   | 12                                 | Placebo capsule  | IADPSG criteria |

BMI: Body mass index; cfu: Colony-forming unit; GDM: Gestational diabetes mellitus; R: Randomized; DB: Double-blinded; PC: Placebo-controlled; OL: Open-label; SB: Single-blinded; GA: Gestational age; IADPSG: The International Association of Diabetes in Pregnancy Study Group; ACOG: The American College of Obstetricians and Gynecologists; NR: Not reported.

Table 2. Study quality evaluation via the Cochrane’s Risk of Bias Tool

| Study           | Random sequence generation | Allocation concealment | Blinding of participants | Blinding of outcome assessment | Incomplete outcome data addressed | Selective reporting | Other sources of bias |
|-----------------|----------------------------|------------------------|--------------------------|--------------------------------|-----------------------------------|---------------------|-----------------------|
| Luoto 2010      | Unclear                    | Unclear                | Low risk                 | Low risk                       | Low risk                          | Low risk            | Low risk              |
| Lindsay 2014    | Low risk                   | Low risk               | Low risk                 | Low risk                       | Low risk                          | Low risk            | Low risk              |
| Wickens 2017    | Low risk                   | Unclear                | Low risk                 | Low risk                       | Low risk                          | Low risk            | Low risk              |
| Okesene 2019    | Low risk                   | Low risk               | Low risk                 | Low risk                       | Low risk                          | Low risk            | Low risk              |
| Pellonpera 2019 | Low risk                   | Low risk               | Low risk                 | Low risk                       | Low risk                          | Low risk            | Low risk              |
| Callaway 2019   | Low risk                   | Low risk               | Low risk                 | Low risk                       | Low risk                          | Low risk            | Low risk              |
| Wang 2019       | Unclear                    | Unclear                | High risk                | Low risk                       | Low risk                          | Low risk            | Low risk              |
| Halkjar 2020    | Low risk                   | Low risk               | Low risk                 | Low risk                       | Low risk                          | Low risk            | Low risk              |
| Cao 2020        | Unclear                    | Unclear                | High risk                | High risk                      | Low risk                          | Low risk            | Low risk              |
| Asgharian 2020  | Low risk                   | Low risk               | Low risk                 | Low risk                       | Low risk                          | Low risk            | Low risk              |
| Shahriari 2021  | Low risk                   | Low risk               | Low risk                 | Low risk                       | Low risk                          | Low risk            | Low risk              |
| Godfrey 2021    | Low risk                   | Unclear                | Low risk                 | Low risk                       | Low risk                          | Low risk            | Low risk              |
| Baloch 2022     | Unclear                    | Unclear                | Low risk                 | Low risk                       | Low risk                          | Low risk            | Low risk              |
| Liu 2022        | Unclear                    | Unclear                | Low risk                 | Low risk                       | Low risk                          | Low risk            | Low risk              |

the remaining 12 studies were double-blinded [26–29, 35–42]. The details of the random sequence generation were reported in nine studies [26, 27, 36–42], and seven studies reported the details of allocation concealment [27, 36, 38–42].

### Influence of probiotics on the incidence of GDM

Pooled results of 14 studies using random-effects models showed that probiotics significantly reduced the incidence of GDM as compared to control (RR: 0.71, 95% CI: 0.52–0.96,  $P = 0.03$ ; Figure 2A) with significant heterogeneity ( $I^2 = 73%$ ). Sensitivity analysis by excluding the study with ACOG criteria [36] for the diagnosis of GDM retrieved similar results (RR: 0.70, 95% CI: 0.51–0.95,  $P = 0.02$ ;  $I^2 = 75%$ ). In addition, a sensitivity analysis excluding the two studies with probiotics started at the 24 weeks of GA [36, 41] also showed similar results (RR: 0.71, 95% CI: 0.51–0.98,  $P = 0.04$ ;  $I^2 = 76%$ ).

The meta-regression showed that the females’ mean BMI was positively associated with the RR for the effect of probiotics on GDM (coefficient = 0.084,  $P = 0.01$ ; Figure 2B and Table 3), which largely explained the source of between-study heterogeneity (residual  $I^2 = 10.5%$ ). Other variables such as sample size, mean age, probiotics dose, median GA for starting probiotics, or incidence of GDM in control groups were not suggested to be significant modifiers for the effect of probiotics on GDM, according to the results of the meta-regression analyses ( $P$  all > 0.05, Table 3).

Subsequent subgroup analyses according to the study country did not significantly affect the results ( $P$  for subgroup difference = 0.09; Figure 3A). However, the results of subgroup analyses indicated that the preventative efficacy of probiotics on GDM was remarkable in females < 30 years, but not in those  $\geq 30$  years (RR: 0.42 vs 1.05,  $P$  for subgroup difference < 0.001; Figure 3B). In addition, it was also indicated

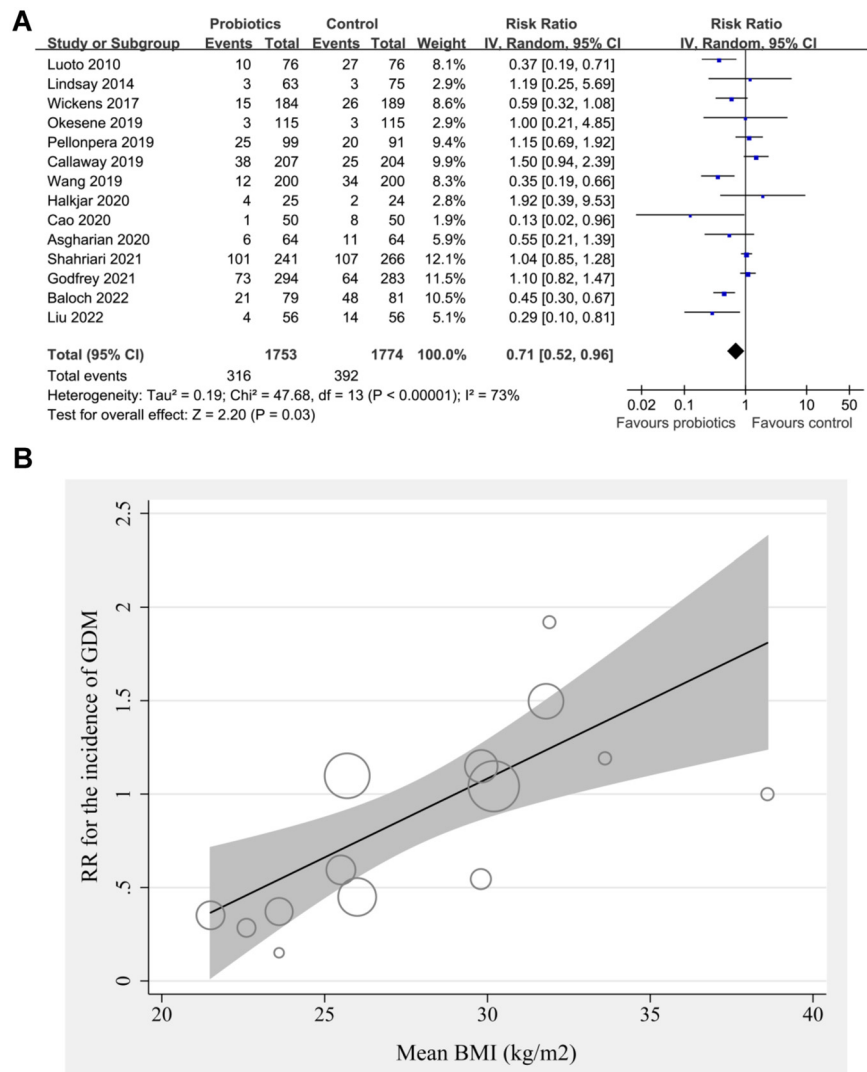
that probiotics significantly reduced the risk of GDM in females with BMI < 26 kg/m<sup>2</sup>, but not in those with BMI  $\geq 26$  kg/m<sup>2</sup> ( $P$  for subgroup difference = 0.001; Figure 4A). Subgroup analyses did not support that other study characteristics could significantly influence the effect of probiotics supplementation on the risk of GDM, such as probiotics dose ( $P$  for subgroup difference = 0.70; Figure 4B), timing of probiotic supplementation ( $P$  for subgroup difference = 0.53; Figure 5A), or risk of GDM as reflected by the incidence of GDM in controls ( $P$  for subgroup difference = 0.97; Figure 5B).

### Publication bias

The funnel plots for the meta-analysis of the probiotics’ influence on the incidence of GDM in pregnant women are shown in Figure 6. The funnel plots are symmetrical on visual inspection, suggesting the low risk of publication bias. The results of Egger’s regression test also suggested a low risk of publication bias ( $P = 0.39$ ).

### Discussion

In our study, by pooling the results of 14 RCTs, we found that probiotics supplementation during pregnancy could significantly reduce the incidence of GDM. Interestingly, subsequent meta-regression and subgroup analyses suggested that the BMI of the pregnant females may significantly modify the effect of probiotics on GDM, which largely explained the source of heterogeneity. Specifically, probiotics significantly reduced the risk of GDM in women with BMI < 26 kg/m<sup>2</sup>, but not in those with BMI  $\geq 26$  kg/m<sup>2</sup>. In addition, the preventative efficacy of probiotics on GDM was remarkable in women < 30 years, but not in those  $\geq 30$  years. Taken together, the results of this meta-analysis indicate that probiotics may be effective in



**Figure 2. Meta-analysis for the role of probiotics on the incidence of GDM in pregnant women.** (A) Forest plots for the overall meta-analysis of the influence of probiotics on the incidence of GDM; (B) Univariate regression analysis for the influence of BMI on the efficacy of probiotics for the prevention of GDM. RR: Risk ratio; CI: Confidence interval; GDM: Gestational diabetes mellitus; BMI: Body mass index; IV: Inverse variance.

**Table 3. Results of univariate meta-regression analysis**

| Variables                                  | RR for the incidence of GDM |                  |          |                         |
|--|-----------------------------|------------------|----------|-------------------------|
|  | Coefficient                 | 95% CI           | P values | I <sup>2</sup> residual |
| Sample size                                | 0.0010                      | 0.0004–0.0024    | 0.15     | 35.2%                   |
| Mean age (years)                           | 0.082                       | –0.081–0.244     | 0.30     | 45.6%                   |
| BMI (kg/m <sup>2</sup> )                   | 0.084                       | 0.025–0.144      | 0.01     | 10.5%                   |
| Dose of probiotics (10 <sup>9</sup> cfu/d) | –0.00057                    | –0.02342–0.02227 | 0.96     | 53.0%                   |
| Median GA for starting probiotics          | 0.0048                      | –0.0398–0.0495   | 0.82     | 53.1%                   |
| Incidence of GDM in control group (%)      | –0.010                      | –0.027–0.007     | 0.23     | 46.3%                   |

RR: Risk ratio; GDM: Gestational diabetes mellitus; CI: Confidence interval; BMI: Body mass index; cfu: Colony-forming unit; GA: Gestational age.

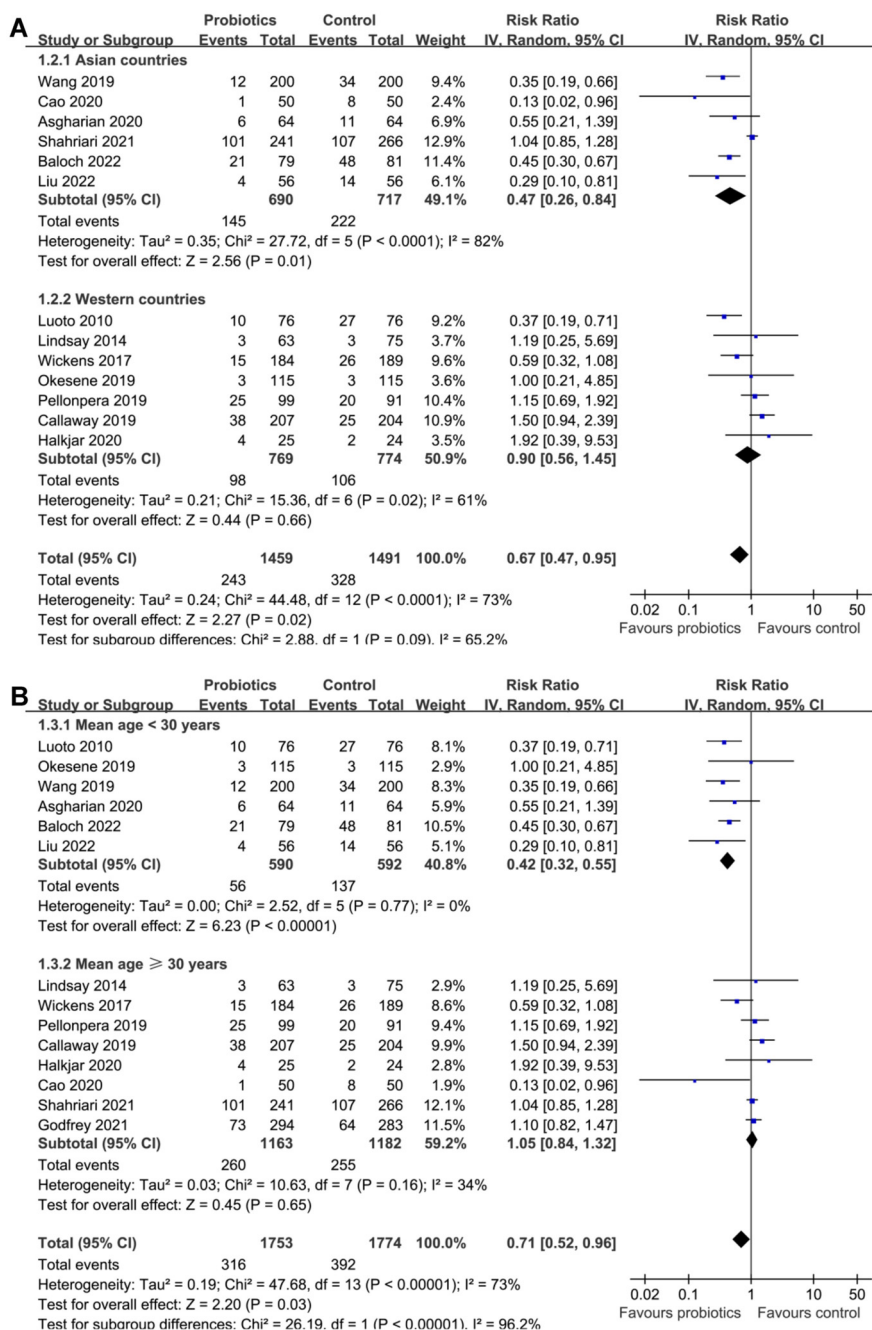
270 reducing the risk of GDM, particularly for females with lower  
 271 BMI and younger age.

272 Although several meta-analyses have been published on the  
 273 topic of the influence of probiotics supplementation on the risk

of GDM [22, 23], this current updated meta-analysis has sev-  
 eral methodological strengths compared to the previous ones.  
 First, in this meta-analysis, we performed an extensive liter-  
 ature search in six commonly used electronic databases, and

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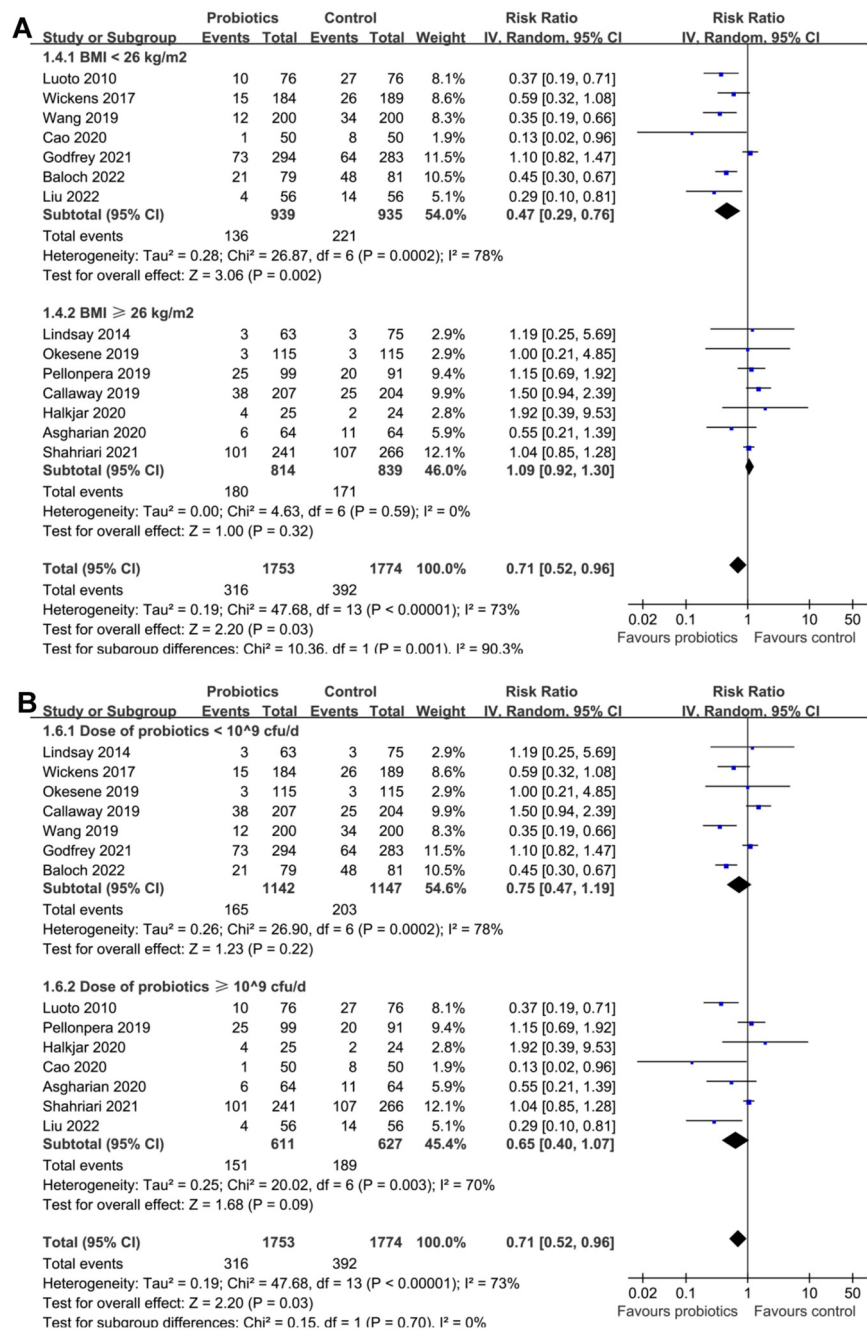


**Figure 3. Forest plots for the subgroup analyses of the efficacy of probiotics for the prevention of GDM.** (A) Subgroup analysis based on country and (B) based on mean ages. GDM: Gestational diabetes mellitus; CI: Confidence interval; IV: Inverse variance.

retrieved eligible RCTs which investigated the efficacy of probiotics for the prevention of GDM. As a result, 14 studies involving 3527 pregnant females were included. The overall sample size of the meta-analysis is much larger than that of the previous ones [22, 23]. In addition, multiple meta-regression and subgroup analyses were performed to identify the study characteristics' influences on the outcome and to determine the source of heterogeneity.

We found that the BMI of the females was positively associated with the RR of the probiotic's effect on GDM, and probiotics significantly reduced the risk of GDM in females with

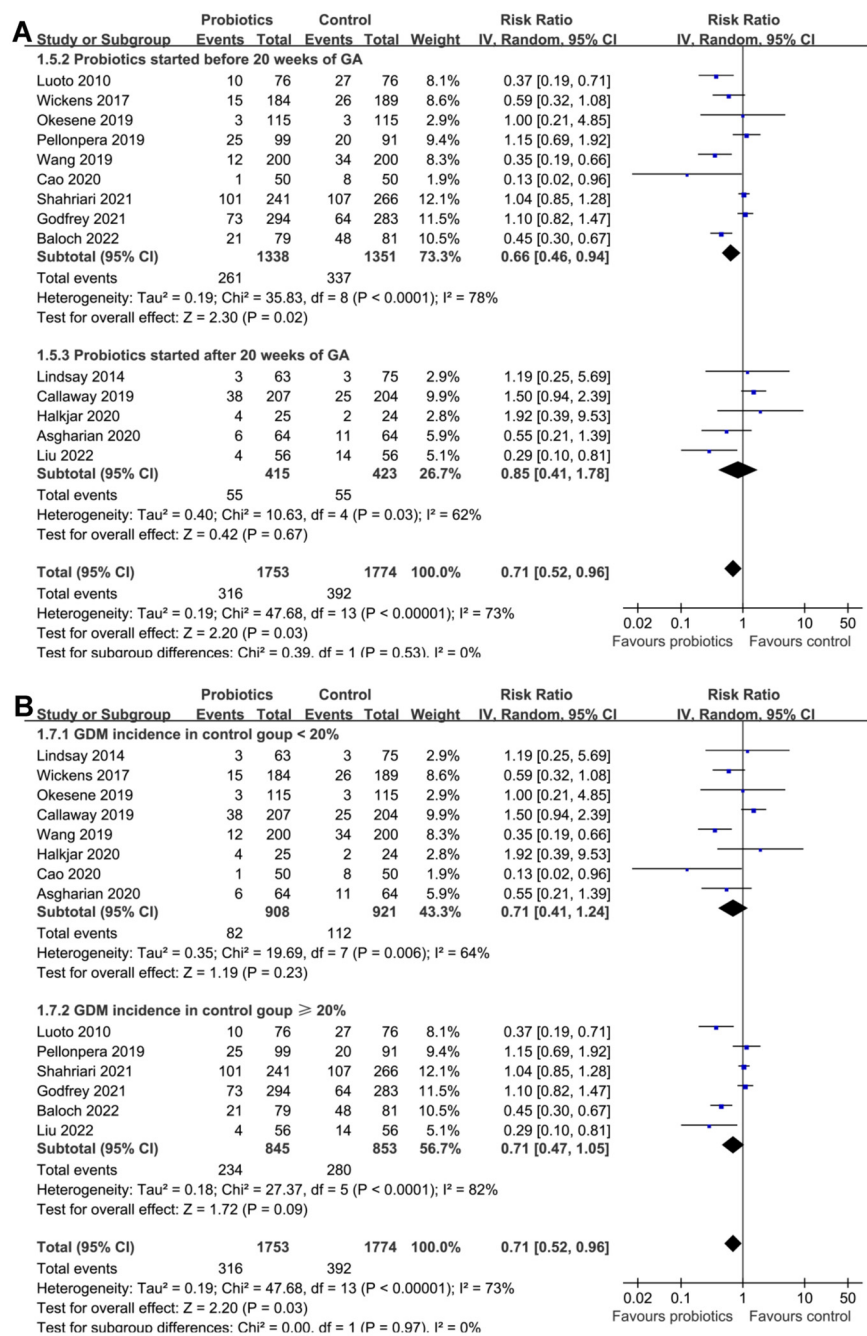
BMI < 26 kg/m<sup>2</sup>, but not in those with BMI ≥ 26 kg/m<sup>2</sup>. Similarly, the preventative efficacy of probiotics on GDM was remarkable in women < 30 years, but not in those ≥ 30 years. The mechanisms underlying these findings remain to be determined. Interestingly, it has been confirmed that advanced maternal age [45] and obesity [46] are established risk factors for GDM. Therefore, the findings of this study may suggest that probiotics supplementation is effective in reducing the risk of GDM in low-risk women, but not in high-risk women. Physiologically, the mechanisms underlying the effects of probiotics supplementation during pregnancy are to attenuate the



**Figure 4. Forest plots for the subgroup analyses of the efficacy of probiotics for the prevention of GDM.** (A) Subgroup analysis based the mean BMI of women and (B) subgroup analysis based on the probiotics dose. GDM: Gestational diabetes mellitus; BMI: Body mass index; CI: Confidence interval; cfu: Colony-forming unit; IV: Inverse variance.

300 gut dysbiosis related to pregnancy, a potential pathway linked  
 301 to the pathogenesis of GDM [47]. For females with high risk  
 302 for GDM, multiple mechanisms may be involved in the patho-  
 303 genesis of GDM besides dysregulation of intestinal microbiota,  
 304 such as islet beta-cell dysfunction, insulin resistance, neurohor-  
 305 monal dysregulation, oxidative stress, and inflammation [48],  
 306 and probiotics supplementation may become less effective. In  
 307 addition, a previous study suggested that multi-strain probi-  
 308 otics are beneficial for improved metabolic and inflammatory  
 309 outcomes in post-GDM women by modulating gut dysbiosis,

which highlighted the necessity for a comprehensive strategy  
 for postpartum treatment that includes probiotics to protect  
 post-GDM women from developing glucose intolerance [49].  
 Accordingly, females with advanced age and obesity may  
 respond poorly to probiotics because they have lesser gut micro-  
 bial diversity. Moreover, a recent meta-analysis indicated that  
 probiotics may have positive effects on metabolic, inflamma-  
 tion, oxidative stress, and neonatal outcomes in females with  
 GDM. Additionally, diet and pre-intervention washout may  
 modify the effects of probiotics [50]. These factors may also



**Figure 5. Forest plots for the subgroup analyses of the efficacy of probiotics for the prevention of GDM.** (A) Subgroup analysis based on the timing of probiotics supplementation and (B) based on the risk of GDM of the included women as reflected by the incidence of GDM in control groups. GDM: Gestational diabetes mellitus; CI: Confidence interval; GA: Gestational age; IV: Inverse variance.

320 confound the efficacy of probiotics in pregnant females with  
 321 high risk for GDM. These hypotheses should be validated in  
 322 future studies.

323 This meta-analysis has limitations. First, the species/strains  
 324 of probiotics varied among the included studies, which may  
 325 also lead to heterogeneity. Future studies should be performed  
 326 to determine the optimal species/strains for the prevention  
 327 of GDM. Second, although our meta-regression and subgroup  
 328 analyses did not show that differences in dose or timing for

starting probiotics supplementation may modify the effect of  
 probiotics for the prevention of GDM, the optimal dose and  
 timing for starting probiotics remain to be clarified in this  
 clinical setting. Third, the incidence of GDM could be signif-  
 icantly affected by dietary habits and physical activities [51],  
 two key factors that may modify the potential preventa-  
 tive efficacy of probiotics on GDM. However, these two fac-  
 tors were rarely reported or controlled among the included  
 studies. In addition, most of the studies did not evaluate the

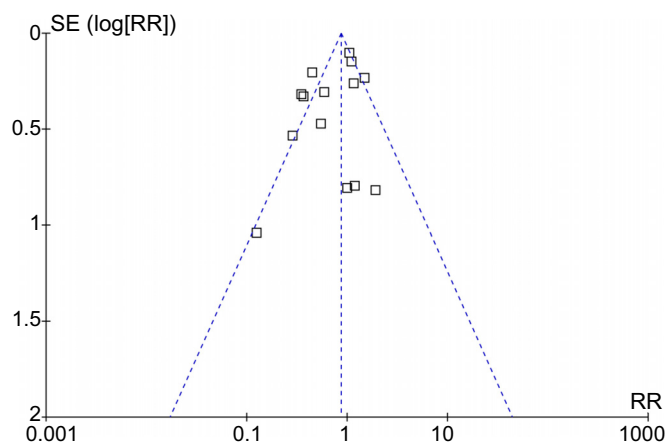


Figure 6. Funnel plots evaluating the publication bias of the meta-analysis for the role of probiotics on the incidence of GDM in pregnant women. GDM: Gestational diabetes mellitus; RR: Risk ratio.

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baseline gut microbial diversity and did not observe the effect of probiotics on gut microbial diversity after intervention. Finally, this meta-analysis was based on study-level data rather than individual patient-level data. Accordingly, results of the meta-regression and subgroup analyses should be interpreted with caution. Large-scale RCTs are still needed to validate these findings.

## Conclusion

Taken together, probiotics supplementation may be effective in reducing the risk of GDM, particularly for females with lower BMI and young age. Although the optimal species/strains, dose, and starting timing of probiotics supplementation remain to be determined, these findings support the potential use of probiotics supplementation as an effective strategy to reduce the incidence of GDM in pregnant females. Further research is needed to evaluate the influence of probiotic supplementation on the risk of GDM in high-risk females, such as those with advanced age and obesity, especially high-quality RCTs.

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